

A NONUNIFORM BOUND FOR THE APPROXIMATION OF POISSON BINOMIAL BY POISSON DISTRIBUTION

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It is well known that Poisson binomial distribution can be approximated by Poisson distribution. In this paper, we give a nonuniform bound of this approximation by using Stein-Chen method.

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1. Introduction and main result. Let X_1, X_2, \dots, X_n be independent, possibly not identically distributed, Bernoulli random variables with $P(X_i = 1) = 1 - P(X_i = 0) = p_i$ and let $S_n = X_1 + X_2 + \dots + X_n$. The sum of this kind is often called a Poisson binomial random variable. In the case where the “success” probabilities are all identical, $p_i = p$, S is the binomial random variable $\mathcal{B}(n, p)$. Let $\lambda = \sum_{i=1}^n p_i$ and let \mathcal{P}_λ be the Poisson random variable with parameter λ , that is, $P(\mathcal{P}_\lambda = \omega) = e^{-\lambda} \lambda^\omega / \omega!$ for all nonnegative integers ω . It has long been known that if p_i 's are small, then the distribution of S_n can be approximated by a distribution of \mathcal{P}_λ (see, e.g., Chen [2]).

In this paper, we investigate the bound of this approximation. As an illustration, we look at the case of $p_1 = p_2 = \dots = p_n = p$. There are at least three known uniform bounds: Kennedy and Quine [6] showed that, for $0 < \lambda \leq 2 - \sqrt{2}$,

$$|P(S_n \leq \omega) - P(\mathcal{P}_\lambda \leq \omega)| \leq 2\lambda[(1-p)^{n-1} - e^{-np}], \quad (1.1)$$

Barbour and Hall [1] showed that

$$|P(S_n \leq \omega) - P(\mathcal{P}_\lambda \leq \omega)| \leq \min(p, \lambda p), \quad (1.2)$$

and Deheuvels and Pfeifer [5] proved that

$$\begin{aligned} & |P(S_n \leq \omega) - P(\mathcal{P}_\lambda \leq \omega)| \\ & \leq \lambda p e^{-\lambda} \left\{ \frac{(np)^{(a-1)}(a-np)}{a!} - \frac{(np)^{(b-1)}(b-np)}{b!} \right\} + R \end{aligned} \quad (1.3)$$

with $a = [np + 1/2 + \sqrt{np + 1/4}]$, $b = [np + 1/2 - \sqrt{np + 1/4}]$, and $|R| \leq (1/2)(2p)^{3/2}/(1 - \sqrt{2p})$, for $0 < p < 1/2$, and $[x]$ is understood to be the integer part of x .

For the general case, Le Cam [7] investigated and showed that

$$\sum_{\omega=0}^{\infty} \left| P(S_n = \omega) - \frac{e^{-\lambda} \lambda^\omega}{\omega!} \right| \leq \frac{16}{\lambda} \sum_{i=1}^n p_i^2. \tag{1.4}$$

It can be observed that the constant $16/\lambda$ will be large when λ is small. Stein [10] used the method of Chen [3] to improve the bound and showed that

$$|P(S_n \leq \omega) - P(\mathcal{P}_\lambda \leq \omega)| \leq (\lambda^{-1} \wedge 1) \sum_{i=1}^n p_i^2 \tag{1.5}$$

for $\omega = 0, 1, 2, \dots, n$ and $\lambda^{-1} \wedge 1 = \min(\lambda^{-1}, 1)$. In case when λ tends to 0, one can see that (1.5) becomes

$$|P(S_n \leq \omega) - P(\mathcal{P}_\lambda \leq \omega)| \leq \sum_{i=1}^n p_i^2. \tag{1.6}$$

In this paper, we consider a nonuniform bound when λ is small, that is, $\lambda \in (0, 1]$ and $\omega \in \{1, 2, \dots, n-1\}$. Note that, when $\omega \notin \{1, 2, \dots, n-1\}$, we can compute the exact probabilities, that is,

$$P(S_n = 0) = \prod_{i=1}^n (1 - p_i), \quad P(S_n = n) = \prod_{j=1}^n p_j, \tag{1.7}$$

$$P(S_n = \omega) = 0, \quad \omega = n + 1, n + 2, \dots$$

In finding the uniform bound, there are several techniques which can be used; for example,

- (i) the operator method initiated in Le Cam [7],
- (ii) the semigroup approach due to Deheuvels and Pfeifer [4],
- (iii) the Chen-Stein technique, see Chen [3] and Stein [10],
- (iv) direct computations as in Kennedy and Quine [6],
- (v) the coupling method, see Serfling [8] and Stein [10].

In the present paper, our argument closely follows the Chen-Stein technique in Chen [3] and Stein [10]. The following theorem is our main result.

THEOREM 1.1. *Let $\lambda \in (0, 1]$ and $\omega_0 \in \{1, 2, \dots, n-1\}$. Then*

$$|P(S_n = \omega_0) - P(\mathcal{P}_\lambda = \omega_0)| \leq \frac{1}{\omega_0} \sum_{i=1}^n p_i^2. \tag{1.8}$$

2. Proof of the main result. Stein [9] gave a new technique to find a bound in the normal approximation to a distribution of a sum of dependent random variables. His technique was free from Fourier methods and relied instead on the elementary differential equation

$$f'(\omega) - \omega f(\omega) = h(\omega) - N(h), \tag{2.1}$$

where h is a function that is used to test convergence and $N(h) = E[h(Z)]$ where Z is the standard normal. Chen [3] applied Stein's ideas in the Poisson setting. Corresponding to the differential equation in the normal case above, one has an analogous difference equation

$$\lambda f(\omega + 1) - \omega f(\omega) = h(\omega) - \mathcal{P}_\lambda(h), \tag{2.2}$$

where $\mathcal{P}_\lambda(h) = E[h(\mathcal{P}_\lambda)]$ and f and h are real-valued functions defined on $\mathbb{Z}^+ \cup \{0\}$. Let $\omega_0 \in \{1, 2, \dots, n - 1\}$ and define $h, h_{\omega_0} : \mathbb{Z}^+ \cup \{0\} \rightarrow \mathbb{R}$ by

$$h(\omega) = \begin{cases} 1, & \text{if } \omega = \omega_0, \\ 0, & \text{if } \omega \neq \omega_0, \end{cases} \quad h_{\omega_0}(\omega) = \begin{cases} 1, & \text{if } \omega \leq \omega_0, \\ 0, & \text{if } \omega > \omega_0. \end{cases} \tag{2.3}$$

Then we see that the solution f of (2.2) can be expressed in the form

$$f_{\omega_0}(\omega) = \begin{cases} \frac{(\omega - 1)!}{\omega_0!} \lambda^{\omega_0 - \omega} \mathcal{P}_\lambda(1 - h_{\omega_0}), & \text{if } \omega_0 < \omega, \\ -\frac{(\omega - 1)!}{\omega_0!} \lambda^{\omega_0 - \omega} \mathcal{P}_\lambda(h_{\omega_0}), & \text{if } \omega_0 \geq \omega > 0, \\ 0, & \text{if } \omega = 0, \end{cases} \tag{2.4}$$

$$\lambda E[f_{\omega_0}(S_n + 1)] - E[S_n f_{\omega_0}(S_n)] = P(S_n = \omega_0) - P(\mathcal{P}_\lambda = \omega_0). \tag{2.5}$$

Let $S_n^{(i)} = S_n - X_i$ for $i = 1, 2, \dots, n$. By using the facts that each X_j takes on values 0 and 1 and that X_j 's are independent, we have

$$\begin{aligned} E[S_n f_{\omega_0}(S_n)] &= \sum_{i=1}^n p_i E[f(S_n^{(i)} + 1)] \\ &= \lambda E[f_{\omega_0}(S_n + 1)] + \sum_{i=1}^n p_i E[f_{\omega_0}(S_n^{(i)} + 1) - f_{\omega_0}(S_n + 1)] \\ &= \lambda E[f_{\omega_0}(S_n + 1)] + \sum_{i=1}^n p_i E\{X_i [f_{\omega_0}(S_n^{(i)} + 1) - f_{\omega_0}(S_n^{(i)} + 2)]\} \\ &= \lambda E[f_{\omega_0}(S_n + 1)] + \sum_{i=1}^n p_i^2 E[f_{\omega_0}(S_n^{(i)} + 1) - f_{\omega_0}(S_n^{(i)} + 2)], \end{aligned} \tag{2.6}$$

which implies, by (2.5), that

$$P(S_n = \omega_0) - P(\mathcal{P}_\lambda = \omega_0) = \sum_{i=1}^n p_i^2 E[f_{\omega_0}(S_n^{(i)} + 2) - f_{\omega_0}(S_n^{(i)} + 1)]. \tag{2.7}$$

From (2.4), it follows that

$$\begin{aligned}
 & f_{\omega_0}(\omega + 2) - f_{\omega_0}(\omega + 1) \\
 &= \begin{cases} -\lambda^{\omega_0 - \omega - 2} \frac{\omega!}{\omega_0!} [(\omega + 1)\mathcal{P}_\lambda(h_{\omega+1}) - \lambda\mathcal{P}_\lambda(h_\omega)], & \text{if } \omega \leq \omega_0 - 2, \\ \lambda^{\omega_0 - \omega - 2} \frac{\omega!}{\omega_0!} [(\omega + 1)\mathcal{P}_\lambda(1 - h_{\omega+1}) + \lambda\mathcal{P}_\lambda(h_\omega)], & \text{if } \omega = \omega_0 - 1, \\ \lambda^{\omega_0 - \omega - 2} \frac{\omega!}{\omega_0!} [(\omega + 1)\mathcal{P}_\lambda(1 - h_{\omega+1}) - \lambda\mathcal{P}_\lambda(1 - h_\omega)], & \text{if } \omega \geq \omega_0. \end{cases}
 \end{aligned}
 \tag{2.8}$$

CASE 1 ($\omega \leq \omega_0 - 2$). Since

$$(\omega + 1)\mathcal{P}_\lambda(h_{\omega+1}) - \lambda\mathcal{P}_\lambda(h_\omega) = e^{-\lambda} \sum_{k=0}^{\omega+1} \frac{\lambda^k}{k!} (\omega + 1 - k),
 \tag{2.9}$$

we have

$$\begin{aligned}
 |f_{\omega_0}(\omega + 2) - f_{\omega_0}(\omega + 1)| &= \lambda^{(\omega_0 - 2) - \omega} \frac{\omega!}{\omega_0!} \left[e^{-\lambda} \sum_{k=0}^{\omega+1} \frac{\lambda^k}{k!} (\omega + 1 - k) \right] \\
 &\leq \frac{(\omega + 1)!}{\omega_0!} \left[e^{-\lambda} \sum_{k=0}^{\omega+1} \frac{\lambda^k}{k!} \right] \\
 &\leq \frac{(\omega_0 - 1)!}{\omega_0!} \\
 &= \frac{1}{\omega_0},
 \end{aligned}
 \tag{2.10}$$

where we have used the facts that $\lambda \in (0, 1]$ and $0 \leq \omega + 1 - k \leq \omega + 1$ in the first inequality and the conditions $\omega \leq \omega_0 - 2$ and $e^{-\lambda} \sum_{k=0}^{\omega+1} (\lambda^k/k!) \leq 1$ in the second inequality.

CASE 2 ($\omega = \omega_0 - 1$). We have

$$\begin{aligned}
 |f_{\omega_0}(\omega + 2) - f_{\omega_0}(\omega + 1)| &= \frac{\lambda^{-1}}{\omega_0} \left[\omega_0 e^{-\lambda} \sum_{k=\omega_0+1}^{\infty} \frac{\lambda^k}{k!} + \lambda e^{-\lambda} \sum_{k=0}^{\omega_0-1} \frac{\lambda^k}{k!} \right] \\
 &\leq \frac{\lambda^{-1}}{\omega_0} \left[e^{-\lambda} \sum_{k=\omega_0+1}^{\infty} k \frac{\lambda^k}{k!} + e^{-\lambda} \sum_{k=0}^{\omega_0-1} (k+1) \frac{\lambda^{k+1}}{(k+1)!} \right] \\
 &= \frac{\lambda^{-1}}{\omega_0} E[\mathcal{P}_\lambda] \\
 &= \frac{1}{\omega_0}.
 \end{aligned}
 \tag{2.11}$$

CASE 3 ($\omega \geq \omega_0$). Since

$$\begin{aligned} & \frac{1\lambda^{\omega+2}}{(\omega+2)!} + \frac{2\lambda^{\omega+3}}{(\omega+3)!} + \frac{3\lambda^{\omega+4}}{(\omega+4)!} + \dots \\ & \leq \lambda^{\omega-\omega_0+2} \left[\frac{\omega_0\lambda^{\omega_0}}{\omega_0!(\omega_0+1)\dots(\omega+2)} + \frac{(\omega_0+1)\lambda^{\omega_0+1}}{(\omega_0+1)!(\omega_0+2)\dots(\omega+3)} + \dots \right] \\ & \leq \frac{\lambda^{\omega-\omega_0+2}}{(\omega_0+1)(\omega_0+2)\dots(\omega+2)} \left[\sum_{k=\omega_0}^{\infty} \frac{k\lambda^k}{k!} \right] \\ & \leq \frac{e^\lambda \lambda^{\omega-\omega_0+2} E[\mathcal{P}_\lambda]}{(\omega_0+1)(\omega_0+2)\dots(\omega+2)} \\ & = \frac{e^\lambda \lambda^{\omega-\omega_0+3}}{(\omega_0+1)(\omega_0+2)\dots(\omega+2)}, \\ & (\omega+1)\mathcal{P}_\lambda(1-h_{\omega+1}) - \lambda\mathcal{P}_\lambda(1-h_\omega) = -e^{-\lambda} \sum_{k=\omega+2}^{\infty} \frac{\lambda^k}{k!} (k - (\omega+1)) < 0, \end{aligned} \tag{2.12}$$

we have

$$\begin{aligned} |f_{\omega_0}(\omega+2) - f_{\omega_0}(\omega+1)| &= \lambda^{\omega_0-\omega-2} \frac{\omega!}{\omega_0!} e^{-\lambda} \sum_{k=\omega+2}^{\infty} \frac{\lambda^k}{k!} (k - (\omega+1)) \\ &\leq \frac{\lambda\omega!}{(\omega+2)!} \leq \frac{1}{(\omega+1)(\omega+2)}. \end{aligned} \tag{2.13}$$

From Cases 1, 2, and 3, we conclude that

$$|f_{\omega_0}(\omega+2) - f_{\omega_0}(\omega+1)| \leq \frac{1}{\omega_0}. \tag{2.14}$$

By (2.7) and (2.14), we have

$$\begin{aligned} & |P(S_n = \omega_0) - P(\mathcal{P}_\lambda = \omega_0)| \\ & \leq \left(\sum_{i=1}^n p_i^2 \right) E[|f_{\omega_0}(S_n^{(i)} + 2) - f_{\omega_0}(S_n^{(i)} + 1)|] \leq \frac{1}{\omega_0} \sum_{i=1}^n p_i^2. \end{aligned} \tag{2.15}$$

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